

## The mathematical challenge

*What is the relative importance of mutation, selection, random drift and population subdivision for standing genetic variation?*

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## Evolution in a spatial continuum

Alison Etheridge

University of Oxford

Joint work with Nick Barton (IST Vienna), Jerome Kelleher (Edinburgh) and Amandine Véber (ENS)

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## The mathematical challenge

*What is the relative importance of mutation, selection, random drift and population subdivision for standing genetic variation?*

*Use pattern of variation in a sample to infer the genealogical relationships between individuals ~> **coalescent models***

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## Other recruits...

Nathanael Berestycki (Cambridge)

Martin Hutzenthaler (Frankfurt)

Tom Kurtz (Madison)

Habib Saadi (Oxford)

Feng Yu (Bristol)

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## The mathematical challenge

*What is the relative importance of mutation, selection, random drift and population subdivision for standing genetic variation?*

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We require **consistent**

- forwards in time models for evolution of population,
- backwards in time models for genealogical trees relating individuals in a sample from the population.

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## The mathematical challenge

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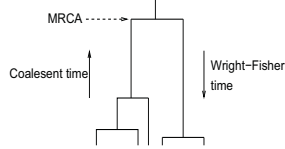
## Drift (large population limit)

Neutral (haploid) panmictic population of constant size

Forwards in time,

- $\mathbb{E}[\Delta p] = 0$  (neutrality)
- $\mathbb{E}[(\Delta p)^2] = \delta t p(1-p)$
- $\mathbb{E}[(\Delta p)^3] = O(\delta t)^2$

Backwards in time



$$dp_t = \sqrt{p_t(1-p_t)}dW_t$$

Coalescence rate  $\binom{k}{2}$ .

$$dp_\tau = \sqrt{\frac{1}{N_e} p_\tau(1-p_\tau)}dW_\tau, \quad \text{Coalescence rate } \frac{1}{N_e} \binom{k}{2}$$

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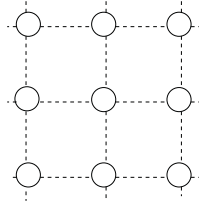
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## Spatial structure

Kimura's stepping stone model

$$dp_i = \sum_j m_{ji}(p_j - p_i)dt + \sqrt{\frac{1}{N_e} p_i(1-p_i)}dW_i$$

System of interacting W-F diffusions



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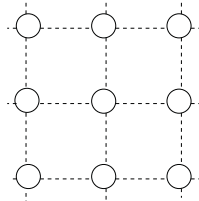
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The coalescent dual process  $\underline{n}$  evolves as follows:

- $\begin{cases} n_i \mapsto n_i - 1 \\ n_j \mapsto n_j + 1 \end{cases}$  at rate  $n_i m_{ji}$
- $n_i \mapsto n_i - 1$  at rate  $\frac{1}{2N_e} n_i (n_i - 1)$



## Drift (large population limit)

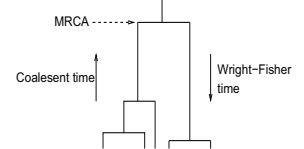
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Coalescence rate  $\binom{k}{2}$ .

## Evolution in a spatial continuum?

For many biological populations it is more natural to consider a spatial continuum. **Malécot** and **Wright** (almost) solved this problem in the 1940s: Assume uniform density, independent reproduction

Identity in state between two genes  $x$  apart:

$$0 = -2\mu F(x) + \frac{1}{2\rho}(1 - F(x))G_{2\sigma^2}(x) + \int [F(y) - F(x)] G_{2\sigma^2}(y - x)dy$$

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$F(0) = \frac{1}{1 + \mathcal{N}/\log(\sqrt{2\mu})}$  where  $\mathcal{N} = 4\pi\rho\sigma^2$  is the *neighbourhood size*.

$F(x) \sim \frac{1}{\mathcal{N}}K_0(|x|/l)$  for  $|x| \gg \sigma$ ,  $l = \sigma/\sqrt{2\mu}$ .

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## Mathematical problems

Felsenstein (1975). **The pain in the torus**: Independent reproduction  
⇒ clumping;

Local regulation ⇒ correlated reproduction.

In 2D the diffusion limit fails over small scales

The obvious backwards model *fails* in 2D

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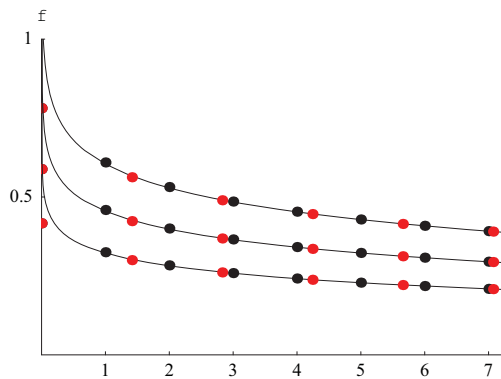
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## ... but



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## Biological problems

Genetic diversity much lower than expected from census numbers

Allele frequencies correlated over long distances

Correlations across loci reflect a **shared** history

*Demographic history of many species dominated by large scale extinction-recolonisation events*

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*Demographic history of many species dominated by large scale extinction-recolonisation events*

... in a spatial continuum, neighbourhood size could be small and then pairwise coalescences may not dominate.

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## $\Lambda$ -coalescents

Pitman (1999), Sagitov (1999)

If there are currently  $n$  ancestral lineages, each transition involving  $j$  of them merging happens at rate

$$\beta_{n,j} = \int_0^1 u^{j-2} (1-u)^{n-j} \Lambda(du)$$

- $\Lambda$  a finite measure on  $[0, 1]$
- Kingman's coalescent,  $\Lambda = \delta_0$

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## An individual based model

- Start with Poisson point process intensity  $\lambda dx$
- At rate  $\mu(dr) \otimes dx \otimes dt$  throw down ball centre  $x$ , radius  $r$ .
- If region empty, do nothing, otherwise:

## Forwards in time

Bertoin & Le Gall (2003)

Suppose there is no Kingman component. The  $\Lambda$ -coalescent describes the genealogy of a sample from a population evolving according to a  $\Lambda$ -Fleming-Viot process.

- Poisson point process intensity  $dt \otimes u^{-2} \Lambda(du)$
- individual sampled at random from population
- proportion  $u$  of population replaced by offspring of chosen individual

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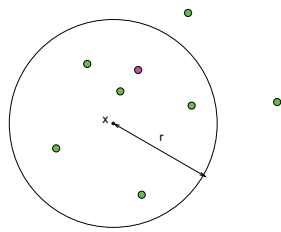
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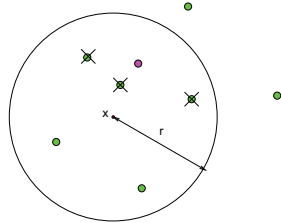
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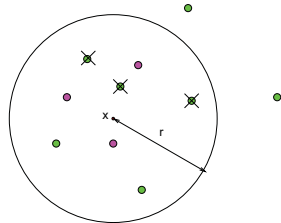
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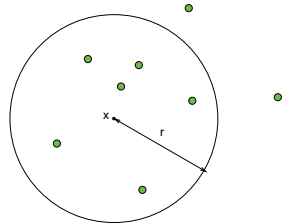


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## A continuum limit

- If  $\lambda$  is sufficiently large, the population survives with positive probability (N. Berestycki, E & Hutzenthaler)

Possibility of empty regions makes formulae cumbersome.

- Replace according to Gaussian density instead of just in disc.
- Let  $\lambda \rightarrow \infty$ . **Model retains signature of finite local population density.**  $\rightsquigarrow$  a *spatial  $\Lambda$ -Fleming-Viot process*

Genealogy of a sample from the population described by a *spatial  $\Lambda$ -coalescent*

Lineages follow coalescing Lévy (actually compound Poisson) processes with *multiple* mergers

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For all  $y \in B_r(x)$ ,

$$\rho(t, y, \cdot) = (1 - u)\rho(t-, y, \cdot) + u\delta_k.$$

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## Mixing

(Amandine Véber)

Work on a torus  $\mathbb{T}(L)$  of side  $L$  in  $\mathbb{R}^2$ . Two types of event:

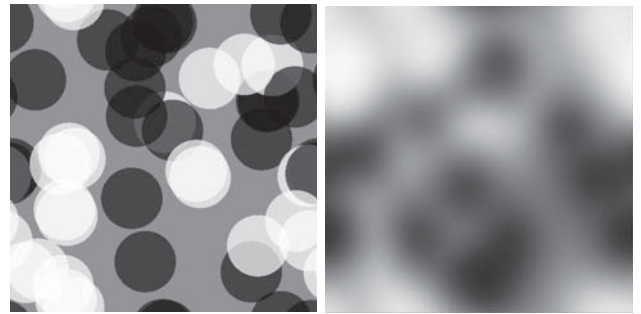
- **Small** events affecting bounded regions.
- **Large** events affecting regions of diameter  $\mathcal{O}(L^\alpha)$

Each ancestral lineage is hit by a *small* event at rate  $\mathcal{O}(1)$ , but by a *large* event at rate  $\mathcal{O}(1/\rho(L))$ .

Sample at random from the whole of  $\mathbb{T}(L)$ .

What happens to the genealogy as  $L \rightarrow \infty$ ?

## Patterns of allele frequencies



## Case (i): $\alpha < 1$

On a suitable timescale the genealogy converges to a Kingman coalescent (with an effective parameter),

The effective population size can depend on both large and small scale events.

c.f. Zähle, Cox, Durrett for classical stepping stone model.

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- $\rho(L) \approx L^2$ , timescale  $\rho(L)$ ,  $\rightsquigarrow$  **spatial**  $\Lambda$ -coalescent in which lineages follow independent Brownian motions in between coalescence events.
- $\rho(L) \approx L^2 \log L$ , timescale  $\rho(L)$ ,  $\rightsquigarrow$  **non-spatial**  $\Lambda$ -coalescent.
- $\rho(L) \gg L^2 \log L$ , timescale  $L^2 \log L$ ,  $\rightsquigarrow$  Kingman coalescent.

*c.f. Nordborg & Krone (2002)*

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## Adding recombination

(Amandine again)

- Small events: Pick **two** parents, types  $ab$  and  $AB$ , say. Write  $r_L$  for fraction of recombinants.

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## Detecting large scale events

Two ideas:

- Slow decay in probability of identity
- Correlations between loci

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## Adding recombination

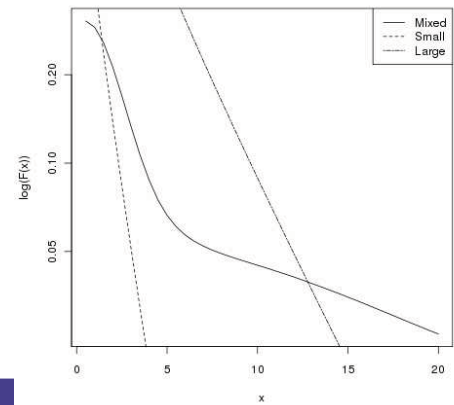
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## Malécot again



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- Large events: ignore recombination.

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## Adding recombination

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## Correlations

Cases... but e.g.  $\rho(L) \leq L^{2\alpha} \rightsquigarrow$

Start  $L^\beta$  apart,  $\beta > \alpha$ .

- If

$$\lim_{L \rightarrow \infty} \frac{\log\left(1 + \frac{\log \rho(L)}{r_L \rho(L)}\right)}{2 \log(L^{\beta-\alpha})} \leq 1$$

then genealogies asymptotically independent.

- Otherwise genealogies completely correlated up to some time  $L^\eta$ .

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- Adding selection. After rescaling one can recover the classical Fisher wave. Small noise perturbations? Spatial hitchhiking. (With Amandine Véber and Feng Yu)

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## A framework for modelling

- Replace  $\mathbb{R}^2$  by an arbitrary Polish space
- Choose a Poisson number of parents at each reproduction event

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## Some work in progress

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